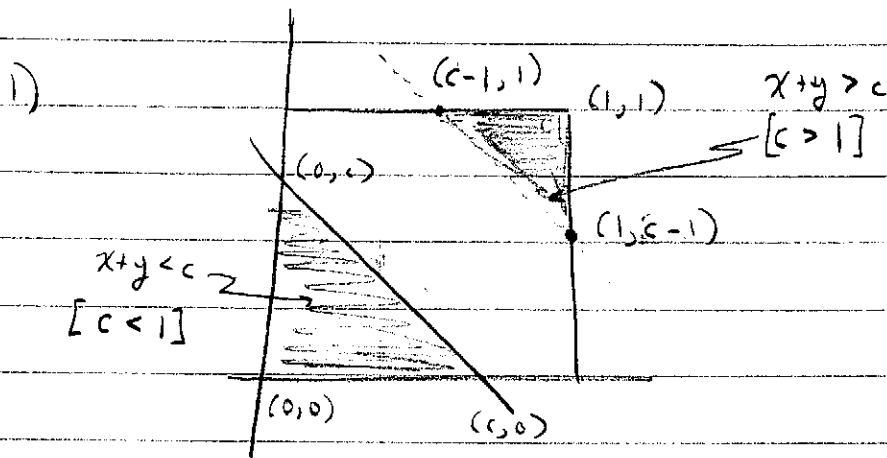


Final Exam Solutions

- 1 -

FCE 603

Fall, 2009



$$S = [0, 2]$$

$A = B$ (restricted to $[0, 2]$, of course)

$$P(W < c) = \begin{cases} c \cdot c / 2, & 0 \leq c \leq 1 \\ 1 - \frac{(2-c)(2-c)}{2}, & 1 \leq c \leq 2 \\ 0, & c < 0 \\ 1, & c \geq 2 \end{cases}$$

$$= \begin{cases} 0, & c \leq 0 \\ c^2 / 2, & 0 < c \leq 1 \\ 2c - 1 - c^2 / 2, & 1 < c \leq 2 \\ 1, & c \geq 2 \end{cases}$$

and

$$P((a, b)) = \begin{cases} (b^2 - a^2) / 2, & 0 \leq a < b \leq 1 \\ 2b - 1 - b^2 / 2 - a^2 / 2, & 0 \leq a \leq 1 < b \leq 2 \\ (2b - 1 - b^2 / 2) - (2a - 1 - a^2 / 2), & 1 < a < b < 2 \\ = 2(b - a) - (b^2 - a^2) / 2 \end{cases}$$

2)

(a)

Consider any $x \in [0, 1]$. Write its binary expansion:

$$0.b_1 b_2 b_3 \dots$$

For any such x , $\exists w = (b_1, b_2, b_3, \dots) \in S$

$\Rightarrow S$ uncountable

(b)

I claim $Y_n \rightarrow 0$ a.s., $\xrightarrow{P}, \xrightarrow{D}$

If $w_i = 0$ for any i , then

$$\lim_{n \rightarrow \infty} Y_n(w) = 0.$$

But the set of w s.t. $w_i = 0$ for some i has $P(A) = 1$. Why? Consider the set A_n such that $w_i = 0$ for some $i \leq n$. Then, $P(A_n) = 1 - 1/2^n$ and $\lim_{n \rightarrow \infty} P(A_n) = 1$.

But it does not go in m.s.

$$E[|Y_n - 0|^2] = \frac{1}{2^n} (1.5)^{2n} + (1 - 1/2^n) \cdot 0 = (2.25/2)^n \rightarrow \infty$$

(c) I claim it converges in distribution to Z

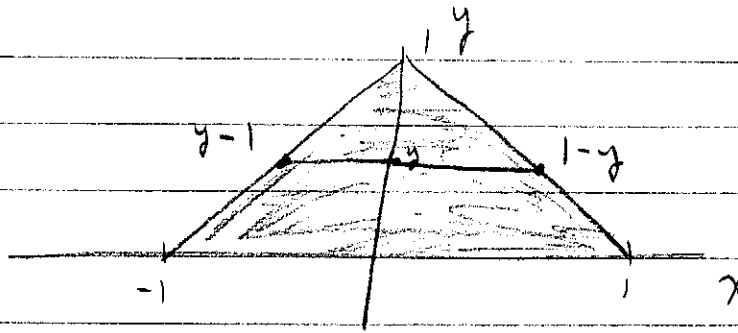
s.t.h. $P(Z=0) = 1/2$ $P(Z=1) = 1/2$ Note that $P(Z_n=0) = 1/2$ $P(Z_n=1) = 1/2$ for all n .

But it does not converge any other way,
in particular; for any

$$P(\{\omega: |X_n(\omega) - X(\omega)| > 1/2\}) \geq 1/2$$

for either n or $n+1$. (for any n); thus,
 X_n does not converge in probability
 $\Rightarrow X_n$ does not converge a.s., m.s.

3)



$$(a) \quad \frac{1}{2} \cdot 2 \cdot 1 \cdot c = 1 \Rightarrow c = 1$$

$$(b) \quad f_Y(y) = \int_{y-1}^{1-y} 1 \, dx = 1-y - (y-1) \\ = 2-2y, \quad 0 \leq y \leq 1$$

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)} \\ = \begin{cases} \frac{1}{2(1-y)}, & 0 \leq y \leq 1 \\ & y-1 \leq x \leq 1-y \\ 0, & \text{else} \end{cases}$$

$$(c) \quad f_{X|Y}(x|1/2) = 1, \quad -1/2 \leq x \leq 1/2$$

$$P(x > 1/4 | Y = 1/2) = \int_{1/4}^{\infty} f_{X|Y}(x|1/2) \, dx \\ = 1/4$$

$$\begin{aligned}
 5) (a) P(Y > 3) &= P(X_1^2 + 2 > 3) \\
 &= P(X_1^2 > 1) \\
 &= P(X_1 \leq -1) + P(X_1 \geq 1) \\
 &= P(X_1 \leq -1) + 1 - P(X_1 \leq 1)
 \end{aligned}$$

$$(X_1 \sim N(0, \sigma^2(0)) = N(0, 1/2))$$

$$\begin{aligned}
 \frac{-1-0}{\sqrt{1/2}} &\stackrel{\downarrow}{=} \frac{1}{2} - \text{erf}(\sqrt{2}) + 1 - \left(\frac{1}{2} + \text{erf}(\sqrt{2})\right) \\
 \frac{1-0}{\sqrt{1/2}} &= 1 - 2 \text{erf}(\sqrt{2})
 \end{aligned}$$

(b) X_2 is a linear combination of jointly Gaussian rvs \Rightarrow Gaussian

$$\mu_{X_2} = E[X_2] = \frac{1}{2} E[Z(0)] + \frac{1}{2} E[Z(1)] = 0$$

$$\begin{aligned}
 \sigma_{X_2}^2 = \text{Var}[X_2] &= E[X_2^2] - \cancel{(E[X_2])^2} \rightarrow 0 \\
 &= \frac{1}{4} E[Z^2(0)] + \frac{1}{2} E[Z(0)Z(1)] + \frac{1}{4} E[Z^2(1)] \\
 &= \frac{1}{4} \cdot \frac{1}{2} + \frac{1}{2} \cdot \frac{1}{2} e^{-1/4} + \frac{1}{4} \cdot \frac{1}{2} \\
 &= \frac{1}{4} + \frac{1}{4} e^{-1/4}
 \end{aligned}$$

$$f_{X_2}(x) = \frac{1}{\sqrt{2\pi\sigma_{X_2}^2}} e^{-x^2/2\sigma_{X_2}^2}$$

$$\begin{aligned}
 P(X_2 > 1) &= 1 - P(X_2 \leq 1) = 1 - \text{erf}\left(\frac{1}{\sqrt{1/4 + 1/4 e^{-1/4}}}\right) \\
 &= 1 - \text{erf}\left(\frac{2}{\sqrt{1 + e^{-1/4}}}\right)
 \end{aligned}$$