

Final Exam Solutions

-1-

FCE 603

Fall, 2002

1)

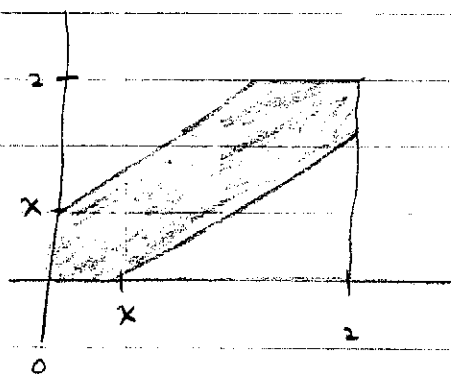
The distance is between 0 and 2. Thus,

$$\Omega = [0, 2]$$

and

\mathcal{A} = Borel set restricted to $[0, 2]$

Need to find distribution (or density function)



$$P(X \leq x) = \frac{\text{Area of strip}}{\text{Total Area}} = \frac{4 - (2-x)^2}{4} = \frac{4x - x^2}{4} = x - \frac{1}{4}x^2$$

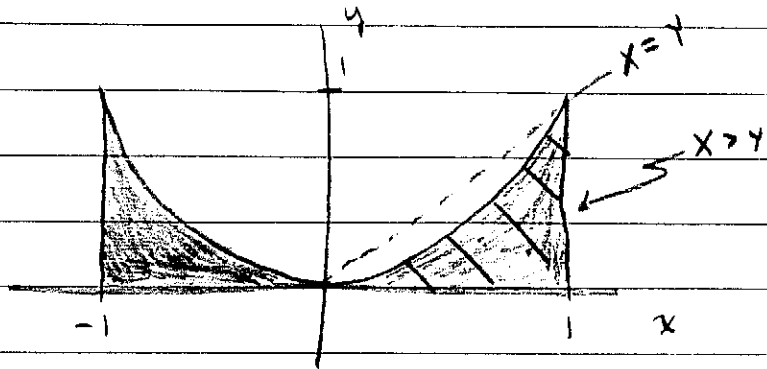
$$\text{Thus, } F_X(x) = \begin{cases} x - \frac{1}{4}x^2, & 0 \leq x \leq 2 \\ 0, & \text{else} \end{cases}$$

and

$$\begin{aligned} P((a,b)) &= b - \frac{1}{4}b^2 - (a - \frac{1}{4}a^2), \quad 0 \leq a \leq b \leq 2 \\ &= (b-a) - \frac{1}{4}(b^2 - a^2) \end{aligned}$$

All other sets in \mathcal{A} are generated from above.

2)



(a)

$$\begin{aligned} \int_{-1}^1 \int_0^{x^2} cy \, dy \, dx &= \int_{-1}^1 \left. \frac{1}{2} cy^2 \right|_0^{x^2} dx \\ &= \int_{-1}^1 \frac{1}{2} cx^4 \, dx = \left. \frac{1}{10} cx^5 \right|_{-1}^1 \\ &= \frac{1}{5} c \Rightarrow c = 5 \end{aligned}$$

(b)

$$\begin{aligned} f_x(x) &= \int_0^{x^2} 5y \, dy = \left. \frac{5}{2} y^2 \right|_0^{x^2} = \frac{5}{2} x^4, \quad -1 \leq x \leq 1 \\ &= \begin{cases} \frac{5}{2} x^4, & -1 \leq x \leq 1 \\ 0, & \text{else} \end{cases} \end{aligned}$$

(c)

$$P(x > y) = \frac{1}{2} \quad (\text{see picture})$$

(Note this requires symmetry about $x=0$
not only of the picture but also the function.)

(d)

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)}$$

$$f_Y(y) = \int_{-1}^{-\sqrt{y}} 5y \, dx + \int_{\sqrt{y}}^1 5y \, dx, \quad 0 \leq y \leq 1$$

$$= 5xy \Big|_{-1}^{-\sqrt{y}} + 5xy \Big|_{\sqrt{y}}^1$$

$$= 5y - 5y^{3/2} + 5 - 5y^{3/2}$$

$$= 10y - 10y^{3/2}$$

$$= \begin{cases} 10y - 10y^{3/2}, & 0 \leq y \leq 1 \\ 0, & \text{else} \end{cases}$$

$$f_{X|Y}(x|y) = \begin{cases} \frac{1}{2(1-y^{1/2})}, & 0 \leq y \leq 1, y \leq x^2 \leq 1 \\ 0, & \text{else} \end{cases}$$

3) $\sigma_x^2 = E[X^2] - (E[X])^2 = 4$

(a) Chebyshev's \neq

$$P(|X - E[X]| \geq \varepsilon) \leq \frac{\sigma_x^2}{\varepsilon^2}$$

$$P(|X - 2| \geq 10) \leq \frac{4}{100} = 1/25$$

(b)

Simply apply erf(.)

$$\begin{aligned} P(X \geq 10) &= 1 - P(X < 12) \\ &= 1 - \left(\frac{1}{2} + \operatorname{erf}\left(\frac{12-2}{2}\right) \right) \\ &= \frac{1}{2} - \operatorname{erf}(5) \end{aligned}$$

4)

Clearly, the limiting random variable is 0 in some sense:

$$P(|X_n - 0| \geq \varepsilon) = 2^{-n} \rightarrow 0 \text{ as } n \rightarrow \infty \text{ for all } \varepsilon > 0.$$

$$\text{Thus } X_n \xrightarrow{P} 0 \Rightarrow X_n \xrightarrow{B} 0$$

Let's try mean square:

$$E[|X_n - 0|^2] = E[X_n^2] = 2^{-n} 2^{2n} = 2^n \rightarrow \infty \text{ as } n \rightarrow \infty$$

Thus, ~~n~~ $\not\xrightarrow{M} 0$

We cannot tell about a.s. since $X_n(\omega)$ is not given explicitly.

5)

(a) No.

$$R_x(\tau) = \rho(\tau) \Leftrightarrow S_x(f) = \text{sinc}(f)$$

but $\text{sinc}(f)$ can be negative, which violates $S_x(f) \geq 0$.

(b)

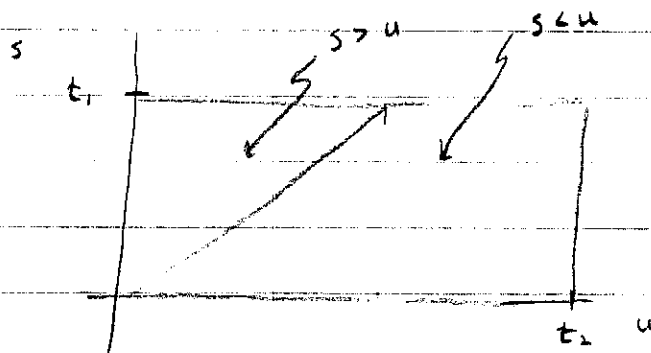
$$R_y(t_1, t_2) = E[Y(t_1)Y(t_2)]$$

$$= E\left[\int_0^{t_1} X(s) ds \int_0^{t_2} X(u) du\right]$$

$$= \int_0^{t_1} \int_0^{t_2} E[X(s)X(u)] du ds$$

$$= \int_0^{t_1} \int_0^{t_2} e^{-3|s-u|} ds du$$

Assume $t_1 < t_2$



$$= 2 \int_0^{t_1} \int_0^s e^{-3(s-u)} du ds + \int_{t_1}^{t_2} \int_0^{t_1} e^{-3(u-s)} ds du$$

$$= 2 \int_0^{t_1} e^{-3s} \left. \frac{1}{3} e^{3u} \right|_0^s ds + \int_{t_1}^{t_2} e^{-3u} \left. \frac{1}{3} e^{3s} \right|_0^{t_1} du$$

$$= \frac{2}{3} \int_0^{t_1} (1 - e^{-3s}) ds + \frac{1}{3} \int_{t_1}^{t_2} (e^{3t_1} e^{-3u} - e^{-3u}) du$$

$$\begin{aligned}
 &= \frac{2}{3} t_1 - \frac{2}{3} \left(-\frac{1}{3} e^{-3s} \right) \Big|_0^{t_1} + \frac{1}{3} \left(e^{3t_1} \left(\frac{1}{3} e^{-3u} \right) \Big|_{t_1}^{t_2} \right. \\
 &\quad \left. + \frac{1}{3} e^{-3u} \Big|_{t_1}^{t_2} \right) \\
 &= \frac{2}{3} t_1 + \frac{2}{9} e^{-3t_1} - \frac{2}{9} - \frac{1}{9} e^{3t_1 - 3t_2} - \frac{1}{9} \\
 &\quad + \frac{1}{3} e^{-3t_2} - \frac{1}{3} e^{-3t_1} \\
 &= \frac{2}{3} t_1 - \frac{1}{9} e^{-3t_1} - \frac{1}{3} + \frac{1}{3} e^{-3t_2} - \frac{1}{9} e^{3(t_1 - t_2)}
 \end{aligned}$$

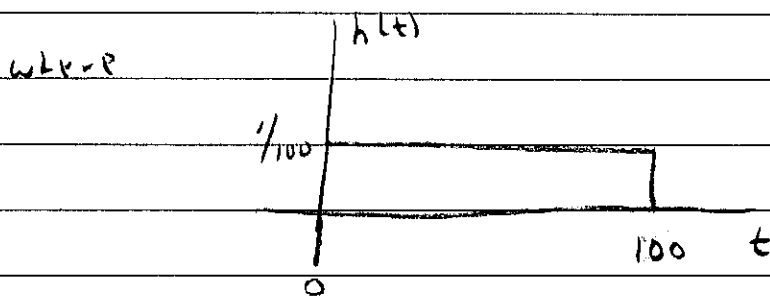
and a similar result holds for $t_2 < t_1$ (with t_1 & t_2 reversed) To show it is not WSS, let $t_1 = t_2$:

$$R_Y(t_1, t_1) = \frac{2}{3} t_1 + \frac{2}{9} e^{-3t_1} - \frac{2}{9}$$

which grows in $t_1 \Rightarrow$ not WSS!

(c)

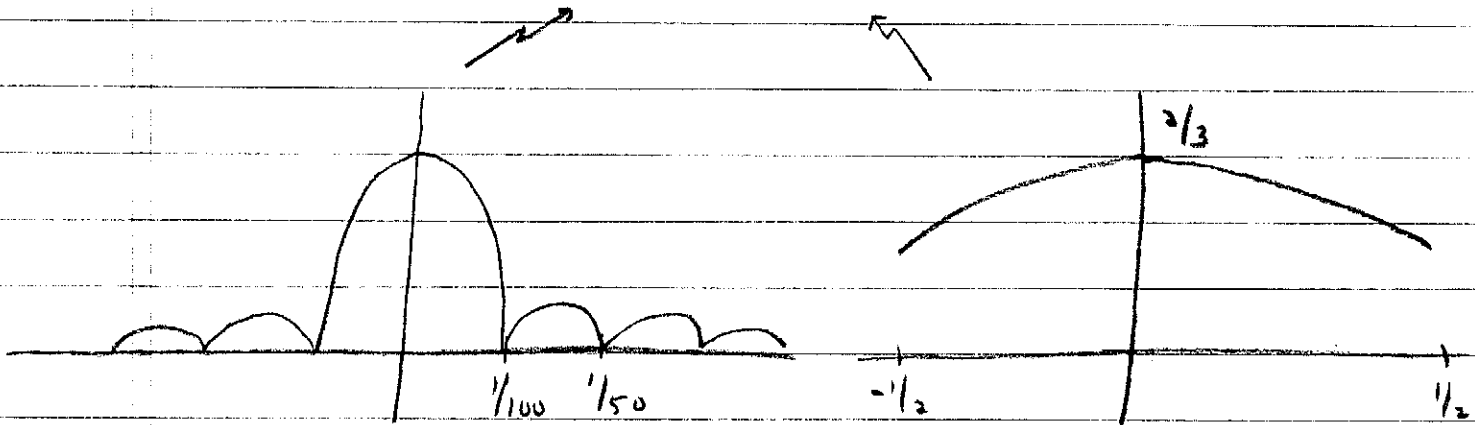
$$S_Y(f) = |H(f)|^2 S_X(f)$$



$$= \frac{1}{100} \text{P} \left(\frac{(t-50)}{100} \right) \longleftrightarrow \text{sinc}(100f) e^{-j2\pi 50f}$$

⇒

$$S_y(f) = \text{sinc}^2(100f) \frac{6}{9 + 4\pi^2 f^2}$$



Thus, $X(t)$ looks almost "white" to this filter, with power spectral density $S_x(f) = 2/3$.

$$\begin{aligned} P_y &\approx \frac{2}{3} \int_{-\infty}^{\infty} \text{sinc}^2(100f) df \\ &= \frac{2}{3} \int_{-\infty}^{\infty} \frac{1}{10^4} p(f/100) df \\ &= \frac{2}{3} \cdot \frac{1}{100} \\ &= \frac{1}{150} \end{aligned}$$

6)

$$\hat{\lambda}_{ML}(Y=y) = \operatorname{argmax}_x P(Y=y | \lambda=x)$$

$$= \operatorname{argmax}_x \frac{e^{-5x} (5x)^{10}}{10!}$$

$$= \operatorname{argmax}_x e^{-5x} (5x)^{10}$$

$$= \operatorname{argmax}_x (-5x + 10 \ln 5x)$$

$$d/dx (-5x + 10 \ln(5x)) = -5 + \frac{10}{5x} \cdot 5 = 0$$

$$x = 2$$

$$\hat{\lambda}_{ML}(Y=10) = 2$$